

Genetic and species-level biodiversity patterns are linked by demography and ecological opportunity

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1 Title: Genetic and species-level biodiversity patterns are linked by

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31 **Abstract:** The processes that give rise to species richness gradients are not well understood, but 32 may be linked to resource-based limits on the number of species a region can support. Ecological 33 limits placed on regional species richness would also limit population sizes, suggesting that these 34 processes could also generate genetic diversity gradients. If true, we might better understand how 35 broad-scale biodiversity patterns are formed by identifying the common causes of genetic 36 diversity and species richness. We develop a hypothetical framework based on the consequences 37 of regional variation in ecological limits to simultaneously explain spatial patterns of species 38 richness and neutral genetic diversity. Repurposing raw genotypic data spanning 38 mammal 39 species sampled across 801 sites in North America, we show that estimates of genome-wide 40 genetic diversity and species richness share spatial structure. Notably, species richness hotspots 41 tend to harbor lower levels of within-species genetic variation. A structural equation model 42 encompassing eco-evolutionary processes related to resource availability, habitat heterogeneity, 43 and human disturbance explained 78% of variation in genetic diversity and 74% of the variation 44 in species richness. These results suggest we can infer broad-scale patterns of species and genetic 45 diversity using two simple environmental measures of resource availability and ecological 46 opportunity.

47

48 Keywords: more individuals hypothesis, heterogeneity, Anthropocene, latitudinal diversity

49 gradient, carrying capacity, macroecology, macrogenetics

50 Introduction

51 Genetic diversity and species richness are the most fundamental levels of biodiversity because 52 they reflect within- and across-species contributions to ecosystem functioning (Oliver et al. 53 2015; Des Roches et al. 2021b). Genetic diversity underlies a population's capacity to adapt in 54 response to environmental change, and species richness enhances ecosystem resiliency to 55 perturbation. If we are to manage the current high rates of biodiversity loss, we need to better 56 understand how patterns of biodiversity are produced and how they interact across levels of 57 biological organization. Patterns of species richness are well-described, but because several 58 independent processes are capable of generating these patterns, their origins remain puzzling. We 59 know less about multi-species patterns of genetic diversity. However, there is good reason to 60 think that the processes forming patterns of species richness could simultaneously produce 61 spatial patterns in neutral genetic diversity (Vellend 2005; Evanno et al. 2009). This is because 62 spatial variation in neutral genetic diversity should reflect how local population-level 63 demographic and evolutionary processes interact with environments to produce species richness 64 gradients. If true, we may be able to infer processes underlying biodiversity patterns at both 65 genetic and species levels by attempting to understand their common causes. The accumulation 66 of open data now allows us to tackle these types of questions by repurposing and synthesizing 67 publicly archived raw data (e.g., Leigh et al. in press; Miraldo et al. 2016; Manel et al. 2020; 68 Schmidt et al. 2020a; Theodoridis et al. 2020; Schmidt and Garroway 2021). Here we produce a 69 continental map of spatial variation in neutral nuclear genetic diversity for North American 70 mammals, show that genetic diversity and species richness covary spatially and are negatively 71 correlated, and find empirical support suggesting that measures of resource availability and

heterogeneity predict both genetic diversity and species richness patterns through their effects ondemography.

74

75 We developed a conceptual framework to explain how genetic diversity and species richness 76 patterns could emerge from common causes. This framework extends predictions from well-77 supported hypotheses for species richness patterns to the population genetic level. Hypotheses 78 for species richness gradients fall into three general categories related to evolutionary time, 79 evolutionary rates, and ecological limits (Mittelbach et al. 2007; Worm and Tittensor 2018; 80 Pontarp et al. 2019). We focus on ecological limits hypotheses—these posit that variation in 81 resource availability limits the number of species able to coexist in a particular area (Rabosky 82 and Hurlbert 2015). Here the speciation, extinction, and colonization dynamics of species are 83 analogous to the birth, death, and immigration dynamics that set carrying capacities at the 84 population level. Simulations suggest multiple hypotheses can produce species richness gradients 85 (Etienne et al. 2019), but the preponderance of theory suggests that ecological limits produce the 86 strongest and most stable gradients (Vellend 2005; Worm and Tittensor 2018; Etienne et al. 87 2019). There is also good empirical support for the likely importance of ecological limits in the 88 formation of species richness patterns (reviewed in Rabosky and Hurlbert 2015; Brodie 2019). 89 We thus considered ecological limits hypotheses as parsimonious starting expectations when 90 exploring the causes of biodiversity patterns (Etienne et al. 2019). 91

92 It is relatively straightforward to extend the consequences of ecological limits on community size
93 to the population genetic level. If environments limit the number of supportable species, they
94 must also limit the population sizes of species, and therefore affect the strength of genetic drift.

95 The first ecological limits hypothesis we consider is the more individuals hypothesis (Wright 96 1983). In terms of community composition, the more individuals hypothesis suggests that 97 resource availability imposes an upper limit on the number of individuals, and as a consequence, 98 the number of species an area can support (Currie 1991; Rabosky and Hurlbert 2015; Storch et 99 al. 2018). Diversity tends to increase with the number of individuals in an assemblage both in 100 terms of genetic diversity within populations and the number of species in a community (Kimura 101 1983; Hubbell 2001). Thus, the more individuals hypothesis predicts neutral genetic diversity 102 and species richness will be positively correlated and increase with resource availability (Fig. 1). 103 The second ecological limits hypothesis we consider pertains to environmental heterogeneity, 104 which includes variation in resources, habitat types, and habitat complexity (Stein et al. 2014). 105 Here we assume heterogeneity equates to niche availability. The idea of area-heterogeneity 106 trade-offs suggests that heterogeneous areas can support richer communities of more specialized 107 species, but these species should tend to have smaller population sizes because resources and 108 species are divided among niches (Kadmon and Allouche 2007; Allouche et al. 2012). Local 109 adaptation, and subsequently specialization, can also occur within and across species distributed 110 across heterogeneous environments. As increasingly specialized populations diverge, genetic 111 variation would be partitioned among locally adapted populations that may eventually no longer 112 interbreed. Compared to larger populations, these smaller populations would also more rapidly 113 lose genetic diversity due to genetic drift. If this were the case, we expect environmental 114 heterogeneity would be positively associated with species richness and negatively associated 115 with neutral genetic diversity (Fig. 1).

Contemporary rapid environmental change also affects biodiversity patterns, yet it is nottypically modelled in a way that makes it comparable to historical processes acting over long

118 periods. A major contemporary ecological limit on diversity is human land transformation. 119 Human activities such as urbanization reduce the amount of habitat available to wild populations 120 (McKinney 2006; Grimm et al. 2008) with consequences at genetic and species levels (Ceballos 121 et al. 2015; WWF 2018; Leigh et al. 2019; Schmidt et al. 2020a). Habitat loss, fragmentation, 122 and homogenization resulting from human land use alters resource and niche availability, thus 123 processes associated with ecological limits should play out in populations and communities of 124 urban wildlife. By reducing habitable area and resource heterogeneity, we predicted that the 125 effects of urbanization for mammals should also cause species richness and genetic diversity to 126 decrease in more heavily disturbed areas (Fig. 1). 127 128 The effects of resource availability and heterogeneity are not mutually exclusive, and in our 129 framework they can act in concert to produce biodiversity patterns. The links among our 130 hypotheses and their predictions are diagrammed in full in Figure 1. We jointly model both 131 hypotheses with a method that allows us to assess their relative importance for shaping genetic 132 diversity and species richness. Our predictions for the ways resource availability and 133 heterogeneity interact are consistent with previous work on species richness in North American mammals (Kerr and Packer 1997), where heterogeneity becomes a more important determinant 134 135 of species richness as resource availability increases. If our model successfully captures known 136 relationships between species richness and environments, and genetic diversity behaves in the 137 ways we predict, we will have strong empirical evidence supporting the contention that 138 continental patterns of neutral genetic diversity and species richness are both are in part governed 139 by ecological limits.

140

141 Our specific objectives were threefold. Because biogeographic patterns of neutral nuclear genetic 142 diversity have not yet been mapped, we first produced a continental map of spatial patterns of 143 genetic diversity in North American mammals. To do this we repurposed publicly archived, raw, 144 neutral nuclear genetic data spanning 38 species and >34,000 individuals at 801 sample sites in 145 the United States and Canada. We then tested the degree to which patterns of genetic diversity 146 matched those of species richness. Having established shared patterns of spatial variation, we 147 then tested our proposed conceptual model based on ecological limits hypotheses where genetic 148 diversity and species richness are caused by common environmental factors (Fig. 1). We tested 149 our hypothetical model using structural equation modelling (SEM), a modelling framework that 150 fits hypothesis networks by accommodating multiple predictor and response variables. Our 151 approach (Fig. 2) allowed us to assess the relative importance of both hypotheses and the effects 152 of contemporary environmental change while accounting for species-level variation using 153 hierarchical models.

154

155 Methods

156 **Biodiversity data**

Genetic diversity. We used raw genotypic data compiled by Schmidt et al. (2020a,b). This data set is comprised of repurposed raw microsatellite data from 34,841 individuals from 38 mammalian species sampled at 801 sites in the United States and Canada. With it we could consistently calculate measures of gene diversity (Nei 1973) and population-specific F_{ST} across sites (Weir and Goudet 2017). See Table 1 for a summary of the dataset. Microsatellite markers estimate genome-wide diversity well (e.g., microsatellite and genome-wide diversity are correlated at $R^2 \sim 0.83$; Mittell et al. 2015). They are commonly used in wildlife population

164	genetic studies because they are cost-effective and do not require a reference genome, which
165	allowed us to maximize sample size. Detailed methods for assembling this dataset can be found
166	in (Schmidt et al. 2020a). Briefly, we performed a systematic search for species names of native
167	North American mammals with keywords "microsat", "single tandem", "short tandem", and
168	"str" using the dataone R package, which interfaces with the DataONE platform to search online
169	open data repositories (Jones et al. 2017). We discarded search results that did not meet our
170	criteria for inclusion and removed results where study design may have influenced genetic
171	diversity. For example, we excluded non-neutral data and samples taken after a recent
172	bottleneck, translocations, managed or captive populations, or island populations. We
173	additionally removed populations with fewer than 5 individuals sampled. Gene diversity
174	estimates the richness and evenness of alleles in a population, and we used it here as our metric
175	for genetic diversity because it is minimally affected by sample size (Charlesworth and
176	Charlesworth 2010) (Fig. S1). Sample sites are treated as point locations.
177	Species richness. We downloaded range maps for terrestrial mammals native to North America
178	from the IUCN Red List database (IUCN 2019). We filtered these maps to retain ranges for
179	extant, native, resident, mainland species in R version 4.0.1 (R Core Team 2020). To generate a
180	map of species richness coincident with genetic sample sites, we estimated species richness by
181	summing the number of ranges overlapping each site.

182

183 Maps and spatial variation partitioning

184 *Genetic diversity and species richness maps.* Our first step was to map spatial patterns in genetic

185 diversity. We accomplished this using distance-based Moran's eigenvector maps (MEMs) in the

186 package adespatial (Dray et al. 2017). MEMs detect spatial patterns in data using a matrix of

187 distances between sites—a neighbor matrix—whose eigenvalues are proportional to Moran's I 188 index of spatial autocorrelation (Borcard and Legendre 2002; Borcard et al. 2004; Dray et al. 189 2006). MEMs are spatial eigenvectors that represent relationships between sites at all spatial 190 scales detectable by the sampling scheme. Multiple MEMs can be included in linear models to 191 identify spatial patterns in data because they are orthogonal. They are appropriate for use in 192 genetics because Moran's I is a direct analog of Malécot's estimator of spatial autocorrelation of 193 allele frequencies (Malécot 1955; Epperson 2005) which accurately summarizes neutral variation 194 in gene flow and allele frequencies (e.g., Sokal and Oden 1978; Epperson 2005). Distance-based 195 MEM analysis produces n-1 MEMs (n being the number of sample sites), but only eigenvectors 196 corresponding to positive spatial autocorrelation are used. MEMs are ordered according to 197 spatial scale explained, with the first eigenvector explaining the broadest autocorrelation pattern. 198 We used linear regressions and the forward selection procedure described in (Blanchet et al. 199 2008) to select two sets of MEMs: one describing spatial patterns in genetic diversity and the 200 other describing species richness. Thirteen MEMs, ranging from broad to fine scales, explained 201 important spatial variation in gene diversity. Forty-three MEMs were important predictors of 202 species richness, and 8 of these patterns were shared by genetic diversity (significant MEMs are 203 listed in Fig. S3).

To restrict ourselves to broad spatial patterns, we focused on genetic and species MEMs with Moran's I values >0.25. We fit individual linear regression models for species richness and genetic diversity with corresponding broad-scale MEMs as covariates and plotted model predicted values representing spatial patterns on maps of North America (Fig. 3). These MEMs describe the broadest-scale spatial patterns at both levels of diversity. By using values of genetic diversity and species richness described by these MEMs, we can visualize pure spatial variation

at the continental scale without local spatial patterns that may be due to environmental

211 idiosyncrasies, and without considering non-spatial variation in genetic diversity. We also

- 212 provided maps of raw genetic diversity and species richness values in Figure S2.
- 213

Variation partitioning. We next quantified the extent to which genetic diversity and species richness covary spatially. Because MEMs for species richness and genetic diversity were computed from the same set of coordinates, they were directly comparable. This allowed us to identify shared spatial MEMs. We used linear regressions and variance partitioning to determine what fraction of the total variation in species richness and genetic diversity could be attributed to: (1) non-spatial variation, (2) non-shared spatial variation, and (3) shared spatial variation. We partitioned variation as follows:

221
$$y_{SR} \sim \alpha + \beta_{1S} (\text{MEM}_{1S}) + \beta_{2S} (\text{MEM}_{2S}) + \dots + \beta_{43S} (\text{MEM}_{43S}) + \epsilon$$

222
$$y_{GD} \sim \alpha + \beta_{1G} (MEM_{1G}) + \beta_{2G} (MEM_{2G}) + \dots + \beta_{13G} (MEM_{13G}) + \epsilon$$

where α is the grand mean, and y_{SR} and y_{GD} are species richness and genetic diversity at sites. MEM_S and MEM_G refer to the set of MEMs explaining spatial variation in species richness (43 MEMs) and genetic diversity (13 MEMs), respectively, and β s are their slopes. The coefficients of variation (R²) for these models give us the proportion of variation in each response variable attributable to spatial variation. Subtracting these values from 1 gives the amount of non-spatial variation.

To determine the amount of shared variation, we used the set of 8 MEMs shared between species
 richness and genetic diversity (MEM_{SG}) as predictors in the regressions below:

231
$$y_{SR} \sim \alpha + \beta_{1SG} (\text{MEM}_{1SG}) + \beta_{2SG} (\text{MEM}_{2SG}) + \dots + \beta_{8SG} (\text{MEM}_{8SG}) + \epsilon$$

232
$$y_{GD} \sim \alpha + \beta_{1SG} (\text{MEM}_{1SG}) + \beta_{2SG} (\text{MEM}_{2SG}) + \dots + \beta_{8SG} (\text{MEM}_{8SG}) + \epsilon$$

R² values from these models estimate the proportion of variation in genetic diversity and species
richness explained by shared spatial variation. Subtracting these values from the total spatial
variation in species richness and genetic diversity gives the proportion of non-shared spatial
variation.

237

238 Structural equation modeling

239 *Population size data.* The more individuals hypothesis is most applicable at broad spatial scales, 240 and when considering the total number of individuals that comprise a species (Storch et al. 241 2018). In place of census sizes for the species in our dataset, which are not consistently available, 242 we craft our hypothesis around species' long-term effective population sizes. The effective 243 population size is a concept defined in population genetics as the number of individuals in an 244 idealized population that experiences the same rate of genetic drift as the real population 245 (Charlesworth and Charlesworth 2010). Populations lose genetic diversity to drift at a rate 246 inversely proportional to the effective population size. Body size is routinely used as a proxy for 247 long-term effective population size at the species level (Frankham 1996; Corbett-Detig et al. 248 2015; Mackintosh et al. 2019; Buffalo 2021). Large bodied species that tend to have long 249 lifespans and produce few offspring generally have smaller effective population sizes than small, 250 fecund, short-lived species (Romiguier et al. 2014; Mackintosh et al. 2019). Thus body size 251 measured at the species level is an imperfect, but nevertheless useful substitution for census size. 252 We recorded adult body mass (g) for each species in our genetic dataset from the PanTHERIA 253 database (Jones et al. 2009) and log-transformed values before analysis. There were no obvious 254 outliers in these data.

255 *Environmental variables.* We used potential evapotranspiration as a surrogate for total ecosystem 256 resource availability (Currie 1991; Rabosky and Hurlbert 2015). Potential evapotranspiration is 257 an indicator of atmospheric energy availability and is one of the strongest environmental 258 correlates of species richness in North American mammals (Currie 1991; Kreft and Jetz 2007; 259 Fisher et al. 2011; Jiménez-Alfaro et al. 2016). As we predict, at the species level, that resource 260 availability across a range sets the long-term effective population size, we estimated mean range-261 wide potential evapotranspiration (mm/yr) using annual data from 1970-2000 available via the 262 CGIAR Consortium for Spatial Information (Trabucco and Zomer 2019). For comparison, we 263 also measured mean range-wide actual evapotranspiration, an alternative measure of resource 264 availability, and present those results in the Supplementary Information. 265 We quantified heterogeneity and niche availability using a 250 m resolution map of land cover 266 types in North America (CEC et al. 2010). This map includes 19 land cover categories based on 267 satellite imagery collected in 2010 with multiple categories of forest, shrubland, grassland, polar 268 habitat types, wetland, cropland, barren land, built up land, and open water. Because the 269 heterogeneity hypothesis suggests species specialize on different resources within their range, we 270 quantified heterogeneity at sites rather than at the species level. We measured heterogeneity 271 using Simpson's diversity index. To assess scale dependence, we calculated Simpson's index 272 within four buffer zones around each site: 5000, 20000, 50000, and 100000 km². Lastly, we 273 quantified human disturbance at each site using human population density (CIESIN 2016) 274 measured within a 10 km buffer following Schmidt et al. (2020a). 275 Analysis. Structural equation modeling accommodates multiple dependent and independent

276 variables in a model network, and directional paths connecting variables represent causal

277 relationships. The strengths of direct paths are regression coefficients (Shipley 2016), and 278 indirect effects can be quantified by multiplying coefficients along paths of direct effects. 279 We constructed a graph of our conceptual model laid out in the introduction (Fig. 1), which we 280 then translated into a network of three linear models for body size, species richness, and genetic 281 diversity. In it, body size is predicted by resource availability, and species richness and genetic 282 diversity are each predicted by body size, heterogeneity, and human disturbance (Fig. 4a). We fit 283 structural equation models using piecewiseSEM in R (Lefcheck 2016; Lefcheck et al. 2019) 284 because this package accommodates complex model structures. We used a linear mixed-effects 285 model with a random intercept for species to account for species-level variation in genetic 286 diversity. PiecewiseSEM fits hierarchical models using the lme4 package (Bates et al. 2015). 287 Body size and species richness models were initially fit as linear regressions, but residuals from 288 both models were spatially autocorrelated at broad scales. We refit these regressions using 289 simultaneous autoregressive models in spatialreg (Bivand et al. 2013) and this successfully 290 removed spatial autocorrelation from the residuals. All variables were scaled and centered before 291 analysis to obtain standardized regression coefficients, allowing us to compare the strength of 292 relationships and the relative support for hypotheses across genetic and species levels.

Our model includes variables measured at the site level (genetic diversity, species richness, heterogeneity, and human disturbance) and species level (body size, resource availability; Fig. 2). This hierarchical data structure introduces spurious correlations between variables sampled at different levels that we know are not causal. For example, regressing human disturbance at sites on species body size would estimate an artefactual experiment (the size of species researchers choose to sample near cities)—not the effects of disturbance on body size. We can account for these known non-causal relationships by allowing variables to have correlated errors (Lefcheck 300 2016). Correlated errors indicate that a relationship exists between variables, but allow the

301 direction of causality to be ambiguous: both could be caused by another factor not included in

302 the model (e.g., researcher species choice). We specified correlated errors between body size and

- 303 human population density, and body size and heterogeneity.
- 304 The conceptual model is evaluated by testing whether additional links are needed between
- 305 variables to make the proposed causal structure more consistent with the data. In piecewiseSEM,
- 306 missing links are tested using tests of directed separation (Shipley 2016), where the null
- 307 hypothesis is that two variables are independent conditional on other predictors in the model. A p
- 308 value for the model network is obtained by testing Fisher's C calculated from the *p* values

309 summed across directed separation tests (Lefcheck 2016; Shipley 2016). A model-wide p < 0.05

310 means the causal structure is not a good fit to the data and additional links are needed to resolve

311 dependencies. If p > 0.05, the model is considered acceptable because we fail to reject our causal

312 structure. This means that although we start with a focus on our conceptual model, the data can

313 suggest the addition or removal of links and our hypotheses can be updated for future testing

314 with new data.

We assessed model fit using R^2 values for each response variable in the model network. For genetic diversity, we used marginal $R^2 (R^2_m)$ which measures the total variation explained by fixed effects, and conditional $R^2 (R^2_c)$ which is the variation explained by both fixed and random effects. For spatial body size and species richness regressions, we report Nagelkerke pseudo- R^2 .

320 Effect of heterogeneity on population divergence

321 After detecting a negative effect of heterogeneity on intraspecific genetic diversity in our
 322 structural equation model, we performed a post hoc analysis to test whether environmental

323 heterogeneity also caused greater population differentiation within species. To test for 324 differentiation we calculated population-specific F_{ST} (Weir and Goudet 2017) as a measure of 325 genetic divergence using the hierfstat package (Goudet and Jombart 2015). Population-specific 326 F_{ST} can be interpreted as a relative estimate of the time since a population has diverged from a 327 common ancestor. This metric required at least 2 sampled populations in the original studies to 328 estimate, and due to this constraint 16 sites were excluded from this analysis. We controlled for 329 isolation-by-distance by including MEMs significantly related to F_{ST} to account for spatial 330 structure. We scaled and centered all variables, then used a linear mixed model controlling for 331 species differences by including it as a random effect.

332

333 Results

334 Spatial patterns in genetic diversity and species richness

335 There was no obvious relationship between latitude and nuclear genetic diversity (Fig. 3). 336 Similar to patterns of species richness, a longitudinal gradient in genetic diversity was the 337 dominant pattern for North American mammals—however, it appears regions with high species 338 richness have lower genetic diversity. We detected spatial patterns at genetic and species levels 339 of diversity. Sixty-five percent of the total variation in species richness and 24% of variation in 340 genetic diversity was spatially structured (Fig. 3). Variance partitioning suggested that 85% of 341 the total spatial variation in genetic diversity, and 32% of spatial variation in species richness 342 was accounted for by spatial patterns shared at both levels of diversity. This shared variation 343 implies that, to an extent, neutral genetic diversity and species richness are simultaneously 344 shaped by spatially structured environmental factors, and these shared processes account for 345 most of the spatial variation in genetic diversity.

346

347 Joint environmental causes of genetic diversity and species richness

We present results from the model using a 5000 km^2 heterogeneity buffer in the main text. 348 349 Results from wider heterogeneity buffers can be found in SI Tables S1-S4. Our conceptual 350 model, updated according to tests of directed separation, fit the data well (SEM p=0.33, Fisher's 351 C= 2.245; Fig. 4, Table S1). Note that for structural equation models, p > 0.05 indicates that we 352 fail to reject our model. There was no residual spatial autocorrelation in body size and species 353 richness model residuals, and genetic diversity model residuals were spatially autocorrelated at 354 local scales (genetic diversity Moran's I = 0.01). These Moran's I values indicate very weak 355 spatial structure in the data, and so we decided not to integrate it into our model. Positive spatial 356 autocorrelation at such short distances is likely an artifact of irregular site locations and the 357 hierarchical nature of the data. A lack of strong spatial autocorrelation in the model residuals 358 suggests that the spatial structure of the diversity data was well captured by our model's 359 environmental covariates (Fig. S3). All predicted links in our conceptual model were supported 360 except that between body size and species richness (Fig. 4). Tests of directed separation 361 suggested additional direct links from resource availability to species richness, and genetic diversity to species richness (Fig. 4). Effects of heterogeneity on genetic diversity were not 362 363 detectable at broader scales (Tables S2-S4). These relationships were consistent using actual 364 evapotranspiration as an alternative measure of resource availability (Table S5).

Resource availability, heterogeneity, and human disturbance, acting both directly and indirectly through species population size, explained 20% of the variation in genetic diversity. The specieslevel variation explained by the random effect brought the total variation in genetic diversity explained by our model to 78%. The same model explained 74% of the variation in species

369	richness. Genetic diversity was strongly negatively related to body size. The direct effects of
370	resource heterogeneity on species richness and genetic diversity were of similar magnitude, and
371	directions of effects were as expected if processes associated with greater resource heterogeneity
372	reduce population sizes, lead to increased genetic drift, and facilitate local adaptation and
373	coexistence (Fig. 4, Table S1). However, because gene diversity is not a measure of divergence,
374	we additionally tested whether environmental heterogeneity predicted evolutionary divergence at
375	the population level. Heterogeneity was positively related to genetic divergence, but the effect
376	was not significant ($\beta = 0.06 \pm 0.04$ SE). Finally, human disturbance negatively affected both
377	species richness and genetic diversity, but its effects were stronger for genetic diversity (Fig. 4,
378	Table S1).
379	
380	Discussion
381	We found striking continental spatial gradients in nuclear genetic diversity, and show that these

patterns are negatively correlated with patterns of species richness in North America (Simpson 1964) (Fig. 3). A considerable portion of the variation in genetic diversity and species richness patterns could be explained by just three environmental factors: resource availability, resource heterogeneity, and human disturbance. This is strong empirical evidence suggesting that genetic diversity and species richness patterns emerge, in part, from the same environmental processes.

Both our maps and our structural equation model suggest that resource availability and
heterogeneity interact to produce biodiversity patterns at genetic and species levels. In North
America, the threshold where environmental heterogeneity presumably becomes a more
important determinant of species richness than resource availability lies roughly along the US-

392 Canada border where potential evapotranspiration reaches ~1000 mm/yr (Kerr and Packer 1997). 393 Near this threshold is also where we see longitudinal patterns of genetic diversity emerge. 394 Although the negative correlation between spatial patterns of genetic diversity and species 395 richness is most apparent in species richness hotspots (particularly in the southwest), structural 396 equation modeling incorporating both hypotheses gives us a more nuanced view of the 397 connections between these patterns. Indeed, effects related to resource availability and 398 heterogeneity were of similar magnitude (Fig. 4). Population size and genetic diversity increased 399 with resource availability, and though the link between species' long-term effective population 400 sizes and species richness was unsupported, species richness also increased with resource 401 availability. Moreover, we detected a positive relationship between genetic diversity and species 402 richness as predicted if population and community size increase with resource availability. It 403 may be that effective population size, measured using species body size, is too coarse an 404 indicator of census population size to detect an effect on species richness at sites—if so, site-405 level measures of genetic diversity could be a better indicator of local population sizes. 406

407 Our results suggest that once a minimum energy threshold is reached, populations can afford to 408 specialize in heterogeneous environments while maintaining viable population sizes. In this way, 409 the interplay between ecological limits and ecological opportunity simultaneously produces 410 biogeographic patterns in genetic diversity and species richness. This interpretation of our results 411 assumes that an environmentally set equilibrium between speciation, immigration and extinction 412 has been reached. There is good evidence for this in North American mammals, where diversification rates have slowed as diversity increased (Alroy 2009; Brodie 2019). However, the 413 414 specific ways environments shape nuclear genetic- and species-level diversity will likely differ

415 across taxa and regions depending on whether or not they have reached equilibrium (e.g., 416 Schmidt et al. 2021). Though we measure contemporary genetic diversity, historical variation in 417 resource availability and heterogeneity likely contribute to the patterns we detect because they 418 reflect whether populations have experienced large contractions in the recent past (Hewitt 2000). 419 However, in the past when communities may not have been at equilibrium, it seems likely that 420 other processes could have been the predominant drivers of biodiversity patterns. Indeed, 421 hypotheses about species richness patterns have likely been a topic of debate for so long because 422 several processes operating with different importance across the timeline of diversification are 423 capable of producing gradients (Etienne et al. 2019). It has been suggested that time for 424 speciation should be most detectable more immediately following broad-scale environmental 425 change, and when all regions are colonized, habitats that provide more opportunities for 426 speciation should over time become the most diverse (Pontarp and Wiens 2017). As diversity 427 increases, diversification rates slow as regions approach equilibrium (Brodie 2019). It follows 428 that the relative importance of evolutionary time and diversification rates as contributors to 429 biodiversity patterns varies with time with patterns ultimately affected by variation in ecological 430 limits (Rabosky and Hurlbert 2015; Pontarp and Wiens 2017; Storch et al. 2018).

431

Contemporary environmental change is our chance to explore pre-equilibrium processes. Cities
are the newest and most rapidly expanding biome, and it is clear that they have already
profoundly affected biodiversity patterns (Palumbi 2001; WWF 2018; Schmidt et al. 2020a). At
this early stage of colonization it is unlikely that urban communities have reached equilibrium,
suggesting processes related to evolutionary time and diversification will predominate until more
niches are occupied. Indeed, there is some evidence that following an initial extinction debt after

438	rapid urbanization, older cities support higher species richness (Aronson et al. 2014). Human
439	disturbance had a negative effect on genetic diversity in our model, and also reduces gene flow
440	in mammals (Schmidt et al. 2020a). This suggests that there is potential for population
441	divergence and local adaptation if new urban niches are exploited and spatially varying selection
442	is sufficiently strong in cities. The extent to which urban populations adapt to local
443	environmental conditions is an ongoing and active field of study, and no consensus has been
444	reached (Lambert et al. 2021). Equilibrium levels of genetic diversity and species richness in
445	urban communities thus seem likely to strongly depend on resource availability and
446	heterogeneity both within and across cities, but these aspects of urban environments are not yet
447	well defined or understood (Norton et al. 2016; Des Roches et al. 2021a).
448	
449	Notably, the negative correlation we find between spatial patterns of species richness and nuclear
450	genetic diversity runs opposite the relatively consistent positive correlations found between
451	species richness and mitochondrial genetic diversity gradients (Martin and McKay 2004; Adams
452	and Hadly 2012; Miraldo et al. 2016; Manel et al. 2020; Theodoridis et al. 2020). Mitochondrial
453	DNA has several idiosyncrasies associated with the specific biology of mitochondria that
454	distinguish it from genetic diversity measured with neutral nuclear DNA (Schmidt and Garroway
455	2021). The most commonly used mitochondrial markers are the protein-coding genes
456	cytochrome oxidase I and cytochrome b, which very likely do not evolve under neutrality
457	(Galtier et al. 2009). Unlike neutral nuclear DNA, mitochondrial genetic diversity is not
458	consistently related to life history, ecological traits, or census and effective population sizes
459	(Bazin et al. 2006; Nabholz et al. 2008; James and Eyre-Walker 2020). Mitochondrial genetic
460	diversity is thus a very different quantity than the neutral nuclear diversity estimates we use here,

and its lack of relationship with population size makes it unsuited for testing hypotheses based
on ecological limits. Using genetic diversity metrics estimated from neutral nuclear DNA allows
us to more directly link environments to species richness through demography, population size,
and by extension, species life history traits which partly set the effective population size.

465

466 Ecosystem sustainability, given environmental perturbations occurring more frequently due to 467 human causes, depends on the resiliency of landscapes, communities, and populations (Oliver et 468 al. 2015). Our framework and the results presented herein suggest that we can understand 469 continental patterns of species richness and genetic diversity using two simple measures of 470 resource availability and heterogeneity. This is potentially informative for conservation practices 471 aiming to conserve both of these levels of biodiversity at once. Maps of neutral nuclear genetic 472 diversity can identify regions where long-term effective population sizes may have been 473 historically small, indicating areas where low levels of neutral genetic diversity are not 474 necessarily of immediate conservation concern (e.g., Yates et al. 2019). However, population 475 declines due to recent human disturbance in areas with historically low genetic diversity may 476 warrant specific attention. Furthermore, designing protected area networks based on species 477 richness to maintain beta diversity, or variation between sites (Bush et al. 2016; Socolar et al. 478 2016), will likely also capture differentiated populations with complementary genetic 479 compositions. The connections between environments, species richness, and genetic diversity we 480 find here suggest we should be able to make informed decisions for the joint conservation of 481 species and genetic diversity with knowledge of few environmental parameters.

482

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498	
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702 **Table 1.** Data summary. Summary of aggregated raw genetic data: mean gene diversity, mean number of loci, median number of

703 individuals at sites per species. Species mass (kg); species richness = mean species richness at sites; energy = mean potential

704 evapotranspiration across species' ranges (mm/yr), heterogeneity = mean land cover diversity (Simpson's Index) within a 5,000 km²

705 zone around a site; human population = mean human population density across sites. Ranges of values are given in parentheses for

706 species with multiple sample sites.

					Species			
Species (# sites)	Gene diversity	# loci	# Individuals	Mass	richness	Energy	Heterogeneity	Human population
Alces alces (2)	0.47 (0.43–0.51)	10	44.5 (40–49)	481	34.5 (34–35)	694.5	0.81 (0.78–0.84)	134.79 (1.45–268.14)
Antilocapra americana (1)	0.67	19	175	46.9	63	1615.39	0.73	3.42
Bison bison (8)	0.47 (0.43–0.51)	29	26.5 (7–31)	620	47.62 (32–69)	745.94	0.68 (0.4–0.82)	1.3 (0.52–3.42)
Canis latrans (41)	0.77 (0.69–0.82)	10	7 (5–10)	12	48.39 (40–55)	1039.97	0.7 (0.27–0.85)	134.92 (1.54–1463.37)
Canis lupus (1)	0.66	12	62	35	44	642.5	0.78	5.85
Glaucomys volans (8)	0.75 (0.65–0.81)	7	18.5 (6–120)	0.07	49.12 (47–51)	1287.09	0.68 (0.58–0.8)	7.31 (1.04–33.59)
Lasionycteris noctivagans (1)	0.83	18	87	0.01	53	1252.72	0.71	2.4
Lasiurus cinereus (1)	0.88	19	132	0.03	53	1126.83	0.71	2.4
Leopardus pardalis (2)	0.47 (0.36–0.58)	10	35 (28–42)	11.9	51 (49–53)	2287.6	0.77 (0.76–0.77)	27.12 (18.73–35.51)
Lepus americanus (39)	0.66 (0.48–0.76)	8	15 (7–100)	1.57	50.44 (2–73)	750.2	0.65 (0.27–0.81)	79.95 (0.69–2711.29)
Lynx canadensis (33)	0.72 (0.46–0.75)	14.15 (14–15)	26 (13–328)	9.77	43.82 (6–50)	655.47	0.73 (0.61–0.81)	4.59 (1.04–56.48)
Lynx rufus (65)	0.73 (0.56–0.79)	14.37 (9–17)	27 (7–141)	6.39	55.6 (33–81)	1459.73	0.62 (0.19–0.83)	189.69 (1.04–3540.4)
Martes americana (29)	0.63 (0.54–0.67)	12	22 (11–47)	0.88	45.86 (42–50)	709.47	0.73 (0.65–0.79)	1.21 (0.69–2.27)
Mephitis mephitis (1)	0.81	9	345	2.4	53	1230.39	0.73	34.19
Microdipodops megacephalus (3)	0.78 (0.73–0.82)	11	62 (49–69)	0.01	64.33 (59–70)	1780.52	0.45 (0.37–0.56)	1.04 (1.04–1.04)
Microdipodops pallidus (2)	0.73 (0.73–0.73)	10	52.5 (42–63)	0.01	59 (58–60)	2045.67	0.25 (0.24–0.27)	1.04 (1.04–1.04)
Myotis lucifugus (65)	0.83 (0.72–0.9)	9.2 (8–11)	33 (11–167)	0.01	44.45 (32–66)	971.21	0.75 (0.21–0.86)	25.96 (0–493.88)
Myotis septentrionalis (15)	0.87 (0.85–0.88)	5	54 (34–110)	0.01	42.13 (41–43)	952.73	0.82 (0.74–0.85)	16.05 (1.04–123.03)
Odocoileus hemionus (67)	0.62 (0.2–0.72)	10.55 (10–18)	29 (7–262)	83.8	57.72 (2–87)	1356.11	0.61 (0.08–0.84)	36.97 (0.83–1213.54)
Odocoileus virginianus (64)	0.81 (0.76–0.84)	14	32.5 (10–79)	75	48.95 (46–52)	1243.86	0.55 (0.32–0.74)	57.42 (6.13–351.21)
Oreamnos americanus (1)	0.52	22	102	71.3	37	743.39	0.79	2.06
Otospermophilus beecheyi (3)	0.75 (0.72–0.78)	11	61 (40–104)	0.6	64.33 (57–69)	1634.31	0.6 (0.54–0.71)	6.18 (3.08–9.03)
Ovis canadensis (16)	0.61 (0.48–0.67)	40.12 (16–210)	42.5 (10–276)	74.6	49.5 (44–63)	1553.5	0.38 (0.15–0.73)	1.27 (1.04–3.42)
Pekania pennanti (34)	0.62 (0.52–0.66)	16	20.5 (7–48)	3.75	45.56 (3–51)	808.85	0.65 (0.25–0.8)	130.82 (0–2620)

Peromyscus leucopus (36)	0.82 (0.75–0.87)	13.19 (10–18)	20 (5–38)	0.02	44.83 (12–53)	1475.67	0.63 (0.3–0.83)	2412.04 (4.86–10000)
Peromyscus maniculatus (10)	0.77 (0.75–0.8)	10.9 (10–11)	12.5 (6–31)	0.02	48.3 (47–51)	1139.04	0.75 (0.62–0.83)	9.44 (1.04–62.41)
Procyon lotor (1)	0.84	10	330	6.37	51	1330.93	0.69	68.91
Puma concolor (13)	0.48 (0.33–0.58)	33.62 (10–46)	51 (21–739)	53.9	61.77 (50–82)	1561.58	0.64 (0.43–0.85)	104.17 (1.04–691.74)
Rangifer tarandus (82)	0.77 (0.45–0.87)	16.96 (14–21)	24 (5–283)	108	40.26 (5–63)	546.54	0.65 (0.28–0.83)	1.11 (0.69–5.11)
Sylvilagus transitionalis (3)	0.42 (0.32–0.47)	10	48 (6–103)	0.81	50.33 (50–51)	1070.02	0.75 (0.7–0.8)	161.12 (43.25–244.49)
Tamiasciurus douglasii (14)	0.65 (0.58–0.72)	9	10.5 (7–24)	0.22	65 (54–70)	1237.95	0.6 (0.53–0.82)	7.76 (1.04–93)
Tamiasciurus hudsonicus (12)	0.66 (0.48–0.77)	9	11 (5–48)	0.2	62.58 (54–68)	804.72	0.61 (0.49–0.73)	1.71 (1.04–5.6)
Taxidea taxus (12)	0.73 (0.43–0.82)	17.33 (12–20)	34.5 (19–649)	7.84	49.25 (39–56)	1403.56	0.59 (0.25–0.81)	44.68 (1.04–297.98)
Ursus americanus (43)	0.72 (0.32–0.82)	15 (8–20)	18 (5–2444)	111	51.35 (0–86)	791.39	0.7 (0.23–0.84)	50.31 (0.52–1065.89)
Ursus arctos (19)	0.67 (0.51–0.77)	9.89 (8–20)	48 (14–729)	196	36.68 (1–68)	568.29	0.6 (0.22–0.75)	1.25 (1.04–5.11)
Ursus maritimus (35)	0.69 (0.56–0.8)	15.09 (8–24)	31 (6–1050)	375	9.17 (2–33)	369.94	0.55 (0.27–0.76)	0.82 (0–1.04)
Vulpes lagopus (3)	0.72 (0.68–0.78)	9	28 (8–42)	3.6	25 (22–27)	455.51	0.73 (0.72–0.75)	1.04 (1.04–1.04)
Vulpes vulpes (16)	0.65 (0.53–0.74)	11.44 (8–13)	30 (9–116)	4.84	44.81 (24–54)	836.89	0.71 (0.39–0.83)	882.65 (1.04–3667.27)



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709 Fig. 1. Framework integrating genetic diversity into ecological limits hypotheses. We focus on

- two major ecological limits pathways which stem from resource availability (the more
- 711 individuals hypothesis) and resource heterogeneity. Under the more individuals hypothesis,
- resource availability across a species range positively affects species' long-term effective
- 713 population sizes (1). Nuclear genetic diversity increases with the effective population size (2). If
- population size is regulated by resource availability, it will be positively associated with
- community size and thus species richness (3). In heterogeneous environments, populations and
- 716 species specialize to different niches but have smaller population sizes (4). Specialization
- reduces gene flow (5) and enhances species' ability to coexist (6). Human land transformation
- reduces wildlife habitat (7) and gene flow among populations (8).

Data synthesis

IABLES	Neutral microsatellite data (e.g., DRYAD)	Species range maps (IUCN)	Body size data (PanTHERIA)
VAR	38 species 801 sites	North American terrestrial mammals	
BIOTIC	Genetic diversity = gene diversity	Species richness = # overlapping ranges	Long-term effective population size = adult mass*
RIABLES	North American Iandcover (CEC)	Population density map (SEDAC)	Potential evapotranspiration (CGIAR-SCI)
V			
BIOTIC	Heterogeneity = land cover diversity (Simpson's Index)	Human disturbance = human population density	Energy availability = range-wide mean potential evapotranspiration*
⊲ An	alysis pipeline		*Units = species
	,		
	Broad spatial pattern in genetic diversity?	Shared spatial variation with species richness?	Common causes of shared patterns?
	Distance-based Moran's eigenvector maps	Spatial variation partitioning	Structural equation modeling
	YES	YES	YES
	Visible longitudinal pattern opposite of species richness; 24% of variation is spatial	85% of spatial variation in genetic diversity shared with species richness	variation in genetic diversity, 74% of variation in species richness

- 721 Fig. 2. Methodological workflow detailing data sources and our series of analyses. For structural
- equation models, variables were either measured at each of 801 sample sites for which genetic
- 723 diversity data was available, or at the species level (n = 38 species).



- 725 Fig. 3. Spatial patterns of biodiversity and environmental factors. (*Top row*) Locations of 801 North American mammal populations
- for which raw microsatellite data was available in public repositories. Point color indicates predicted values of genetic diversity and
- species richness based on spatial patterns detected in the data. The variation partitioning plot shows the proportion of variation in
- genetic diversity and species richness which can be explained by spatial factors. Spatial variation is further broken down into shared

- and non-shared spatial variation. (Bottom row) Major environmental features of North America. Note land cover is categorical and
- 730 colors represent different types. Elevation is shown for reference, but was not included in our models.



Fig. 4. Structural equation models. (a) Our conceptual hypothesis network, modified from Figure 1 to accommodate variables measured at species and site levels. Single-headed arrows represent unidirectional relationships between variables. Grey double-headed arrows indicate variables with correlated errors that were excluded from model evaluation (see Methods). (b) Structural equation model results. Green and black lines are positive and negative relationships, respectively, and the grey line is an unsupported link. Line widths reflect partial regression coefficients, which are listed for each path with standard errors. R^2 values are the amount of variation explained for each response variable. Genetic diversity was fit with a random effect for species: R^2_m is the variation explained by fixed effects only, and R^2_c is the variation explained by fixed and random effects.

Supplementary Information for: Genetic and species-level biodiversity patterns are linked by demography and ecological opportunity

Includes: Figures S1-S3 Tables S1-S5



Figure S1. Plot of gene diversity vs. sample size. Gene diversity as a metric of genetic diversity depends on allele frequencies and is minimally affected by sample size. Larger populations have more rare alleles, which contribute little to gene diversity.



Figure S2. Maps of raw genetic diversity (gene diversity) and species richness data for each site. Clear gradients are visible in species richness but not genetic diversity. This is because most (65%) variation in species richness was spatial, in contrast to genetic diversity where comparatively less (24%) variation was spatially structured (Fig. 3).



Figure S3. Correlation coefficients for spatial patterns (MEMs) and environmental variables measured at the site level: potential evapotranspiration (PET), land cover diversity, and human population density. MEMs describe spatial patterns in genetic diversity, species richness, or both (shared spatial patterns). MEMs are ordered from broad (MEM1) to fine scale (MEM194) patterns. Strong correlations indicate that environmental variables included in structural equation models account for broad scale spatial patterns present in genetic diversity and species richness.

Table S1. Path coefficients and standard errors for structural equation model presented in main

2 text (Fisher's C = 2.25, p = 0.33, 2 degrees of freedom). Heterogeneity is measured within 5000

 km^2 of a site.

Response	Predictor	Estimate ± SE
genetic diversity	human population density	-0.16 ± 0.03
genetic diversity	body size	-0.56 ± 0.15
genetic diversity	heterogeneity	-0.06 ± 0.02
body size	PET	-0.17 ± 0.03
species richness	body size	0.02 ± 0.02
species richness	heterogeneity	0.05 ± 0.02
species richness	human population density	-0.05 ± 0.02
species richness	PET	0.04 ± 0.02
species richness	genetic diversity	0.06 ± 0.02
Correlated errors		Partial correlation coefficient
body size	human population density	0.10*
body size	heterogeneity	-0.02

Table S2. Path coefficients and standard errors for structural equation model; heterogeneity is measured within 20000 km² of a site (Fisher's C = 2.09, p = 0.35, 2 degrees of freedom).

Response Predictor		Estimate ± SE
genetic diversity	human population density	-0.16 ± 0.03
genetic diversity	body size	-0.57 ± 0.15
genetic diversity	heterogeneity	-0.04 ± 0.03
body size	PET	-0.17 ± 0.03
species richness	body size	0.02 ± 0.02
species richness	heterogeneity	0.06 ± 0.02
species richness	human population density	-0.05 ± 0.02
species richness	PET	0.05 ± 0.02
species richness	genetic diversity	0.06 ± 0.02
Correlated errors		Partial correlation coefficient
body size	human population density	0.10*
body size	heterogeneity	-0.03

Table S3. Path coefficients and standard errors for structural equation model; heterogeneity is measured within 50000 km² of a site (Fisher's C = 1.78, p = 0.41, 2 degrees of freedom).

Response	Predictor	Estimate ± SE
genetic diversity	human population density	-0.17 ± 0.03
genetic diversity	body size	-0.57 ± 0.15
genetic diversity	heterogeneity	0.00 ± 0.03
body size	PET	-0.17 ± 0.03
species richness	body size	0.02 ± 0.02
species richness	heterogeneity	0.06 ± 0.02
species richness	human population density	-0.06 ± 0.02
species richness	PET	0.05 ± 0.02
species richness	genetic diversity	0.06 ± 0.02
Correlated errors		Partial correlation coefficient
body size	human population density	0.10*
body size	heterogeneity	-0.03

Table S4. Path coefficients and standard errors for structural equation model; heterogeneity is om).

measured within 100000 km ² of a site (Fisher's C = 1.54, $p = 0.46$, 2 degree	s of freedo
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Response	Predictor	Estimate ± SE
genetic diversity	human population density	-0.18 ± 0.03
genetic diversity	body size	-0.56 ± 0.15
genetic diversity	heterogeneity	0.04 ± 0.03
body size	PET	-0.17 ± 0.03
species richness	body size	0.02 ± 0.02
species richness	heterogeneity	0.05 ± 0.02
species richness	human population density	-0.05 ± 0.02
species richness	РЕТ	0.04 ± 0.02
species richness	genetic diversity	0.06 ± 0.02
Correlated errors		Partial correlation coefficient
body size	human population density	0.10*
body size	heterogeneity	-0.03

Table S5. Path coefficients and standard errors for structural equation model using actual

evapotranspiration (AET) as a measure of resource availability. Heterogeneity is measured

within 5000 km² of a site (Fisher's C = 0.53, p = 0.77, 2 degrees of freedom). Genetic diversity $R_m^2 = 0.20$; $R_c^2 = 0.78$; body size $R^2 = 0.48$; species richness $R^2 = 0.74$.

Response	Predictor	Estimate ± SE
genetic diversity	human population density	-0.16 ± 0.03
genetic diversity	body size	-0.56 ± 0.15
genetic diversity	heterogeneity	-0.06 ± 0.02
body size	AET	-0.33 ± 0.03
species richness	body size	0.03 ± 0.02
species richness	heterogeneity	0.04 ± 0.02
species richness	human population density	-0.05 ± 0.02
species richness	AET	0.05 ± 0.02
species richness	genetic diversity	0.05 ± 0.02
Correlated errors		Partial correlation coefficient
body size	human population density	0.09*
body size	heterogeneity	-0.01