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1 Investigating microbial associations from sequencing survey data with

2 co-correspondence analysis

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- 24
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- 1

26 Abstract

27 Microbial communities, which drive the major ecosystem functions, are 28 composed by a wide range of interacting species. Understanding how microbial 29 communities are structured and the underlying processes is a crucial task for 30 interpreting ecosystem response to global change but it is challenging as microbial 31 interactions cannot usually be directly observed. Multiple efforts are currently focused 32 to combine next-generation sequencing (NGS) techniques with refined statistical 33 analysis (e.g., network analysis, multivariate analysis) to characterize the structures of 34 microbial communities. However, most of these approaches consider a single table of 35 sequencing data measured for several samples. Technological advances now make it 36 possible to collect NGS data on different taxonomic groups simultaneously for the same 37 samples and lead to analyze a pair of tables. Here, an analytic framework based on co-38 correspondence analysis (CoCA) is proposed to study the distributions, assemblages and 39 interactions between two microbial communities. We showed the ability of this 40 approach to highlight the relationships between two microbial communities, using two 41 data sets exhibiting various types of interactions. CoCA identified strong association 42 patterns between autotrophic and heterotrophic microbial eukaryotes assemblages, on 43 one hand, and between microalgae and viruses, on the other hand. We demonstrate also 44 how CoCA can be used, in complement to network analysis, to reorder <u>co-occurrence</u> 45 networks and thus investigate the presence of patterns in ecological networks.

46 Introduction

47 Microbial communities are highly diverse (Rappé & Giovannoni, 2003) and drive 48 the major ecosystem functions (e.g., carbon sequestration, climate regulation, gas 49 regulation, nutrient cycling; Ducklow, 2008; Falkowski, Fenchel, & Delong, 2008; 50 Hutchins & Fu, 2017). Understanding how these systems are structured and identifying 51 the underlying processes is a crucial task to predict communities and ecosystem 52 responses to global change (Fuhrman, 2009; Graham et al., 2016). Biotic interactions 53 across taxonomic groups (e.g., predation, parasitism, mutualism or competition) are of 54 broad interest because they are expected to influence the structure and composition of 55 communities (Wisz et al., 2013). Unfortunately, our understanding of the underlying 56 assemblage rules of microbial communities is still limited (Little et al., 2008; Cordero & 57 Datta, 2016).

58 The emergence of high-throughput sequencing techniques (next-generation 59 sequencing; NGS) gave access to the diversity of whole microbial communities, including the non-cultivable fraction (Handelsman, 2004; Zimmerman, Izard, Klatt, Zhou, & 60 61 Aronson, 2014; Zinger, Gobet, & Pommier, 2011). With the large amount of data 62 generated in a single NGS experiment, powerful statistical methods are needed to assess 63 and explain structural patterns in such complex data sets (Bálint et al., 2016; Paliy & 64 Shankar, 2016). A common approach is to combine NGS techniques with network 65 analysis to represent and characterize interactions between partners in microbial communities (Cardona, Weisenhorn, Henry, & Gilbert, 2016; Vacher et al., 2016). 66 67 Various computational methods have been developed to infer networks from NGS data 68 sets (e.g., CoNet: Faust et al., 2012; SparCC: Friedman & Alm, 2012; REBACCA: Ban, An, & 69 Jiang, 2015; CCLasso: Fang, Huang, Zhao, & Deng, 2015; SPIEC-EASI: Kurtz et al., 2015). 70 Another popular approach is to use ordination methods to extract information from NGS

data and describe the variations in community composition among samples (Paliy &
Shankar, 2016). Ordination methods arrange objects in a multidimensional space using
directly the original raw data table (e.g., principal component analysis, correspondence
analysis), or after computing a distance matrix (e.g., non-metric multidimensional
scaling, principal coordinate analysis) (Legendre & Legendre, 2012).
These different approaches are based on a single table composed of read counts
for each Operational Taxonomic Unit (OTU) measured for several samples.

78 Technological advances now make it possible to acquire NGS data on different 79 taxonomic groups simultaneously for the same samples (Fierer et al., 2007) and lead to 80 analyze a pair of tables (i.e., OTUs composition for the same sampling sites for two 81 different taxonomic groups). To analyze such pair of tables, a common practice consists 82 in merging the two tables into a single one and then applying network analysis (e.g., 83 Kueneman et al., 2016; Banerjee et al., 2016; Ma et al., 2016) and/or multivariate 84 analysis (Cannon et al., 2017; Bergelson, Mittelstrass, & Horton, 2019). However, this 85 data aggregation is unsuitable especially when NGS data sets, which are a function of 86 sequencing depth (Ni, Yan, & Yu, 2013), are standardized by dividing read counts by the 87 total number of reads in each sample. In this case, the normaliszation step and further analysis are very sensitive to the difference in number of OTUs and associated counts in 88 89 each taxonomic group. Hence, it is important to use techniques allowing <u>for the</u> analy<u>sis</u> 90 of a pair of NGS data tables while preserving the original structure of the data. In 91 addition, these approaches must be able to mitigate the statistical bias stemming from 92 high-dimensionality (i.e., a number of samples substantially lower than the number of 93 variables), sparsity (i.e., a high proportion of zero counts), and the compositional nature 94 (i.e., a non-independence of relative abundances induced by the row-sum <u>normalisation</u>) 95 that characterize NGS data (Li, 2015). For network analysis, the SPIEC-EASI method has

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96 been adapted to infer associations among microorganisms in a cross-domain analysis 97 (Tipton et al., 2018). For multivariate analyses, several two-table methods exist and 98 have been presented to microbial ecologists in methodological reviews (Ramette, 2007; 99 Buttigieg & Ramette, 2014; Paliy & Shankar, 2016). However, these works focused on 100 asymmetric methods (e.g., canonical correspondence analysis and redundancy analysis) 101 that aim to explain the composition of microbial communities by a limited number of 102 environmental predictors. Unfortunately, these methods are not adapted to link two 103 NGS data tables as they require that there are fewer predictor variables than samples 104 (Dray, Chessel, & Thioulouse, 2003) and thus are not able to deal with the high-105 dimensionality of NGS data. Moreover, these methods compute linear combinations of 106 the predictor variables, which is not suitable if the table of predictors contains 107 community data that display unimodal structure and/or are sum <u>normalised</u> (as the NGS 108 data).

109 This study aims to propose an analytical framework based on co-correspondence 110 analysis (CoCA; ter Braak & Schaffers, 2004), a two-table coupling method developed in 111 community ecology, to study the distributions and assemblages between two microbial 112 communities. This framework is based on correspondence analysis, a method that 113 effectively handles proportional data that contain many zeroes (Gauch et al., 1977; 114 Jackson, 1997; Greenacre, 2009), like NGS data (Paulson, Stine, Bravo, & Pop et al., 115 2013). We show how this method allows to extract information about the co-structure 116 among two microbial communities to estimate the congruence between them. Finally, 117 we show that the outputs of the method can be used to reorder <u>co-occurrence</u> networks 118 inferred by network analysis to enhance the visualization of microbial association and 119 the understanding of assemblage patterns within networks. Hence, our approach echoes 120 to the current lively debate about the practices to create network visualizations which

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121 are both aesthetically appealing and have high information content (see Pocock et al.,

122 2016 for review). We illustrate our approach using two real data sets, one on

123 autotrophic and heterotrophic microbial eukaryotes in shallow freshwater systems and

124 another on microalgae and viruses in marine systems.

125

126 Materials and methods

127 Studying cross-taxon congruence by co-correspondence analysis

Co-correspondence analysis (ter Braak & Schaffers, 2004) is part of the class of 128 canonical analyses with the feature to be designed to analyze a pair of tables containing 129 130 abundance data and to study the co-variations between two types of communities (e.g., 131 plants and pollinators). CoCA is based on correspondence analysis (Benzécri, 1969; Hill, 132 1973) and preserves its fundamental properties of weighted averaging and the use of the chi-square (χ^2) distance for both tables <u>(Supporting Information</u>). The χ^2 distance is 133 134 particularly adapted to NGS data as it handles properly zero values (and in particular 135 double absences) and thus is not hampered by zero inflation (Legendre & Legendre, 136 2012).

Here, co-correspondence analysis in its predictive form (pCoCA) is used for the 137 microalgae-virus data set while the symmetric form (sCoCA) is used for the microbial 138 eukaryote data set. Let $X = [x_{ii}]$ and $Y = [y_{ik}]$ be $n \times p$ and $n \times q$ tables containing the 139 140 relative abundances of each p and q species of two communities measured at the same n141 samples. In the first case, pCoCA is chosen because we wish to investigate the 142 composition of virus communities <u>under the hypothesis that the occurrence of viruses</u> depends mainly on whether microalgal hosts are present or absent at a particular 143 144 sample location. In the second case, sCoCA is chosen as we simply wished to study the 145 relationships between autotrophic and heterotrophic microbial eukaryotes. In its

146 predictive form, CoCA is based on partial least-squares regression analysis (PLS) in a 147 SIMPLS version (ter Braak & Schaffers, 2004) to deal with high-dimensionality and 148 subsequent collinearity in the table of explanatory variables. PLS searches for a linear 149 regression model from a set of orthogonal components (called latent factors) built from 150 collinear explanatory variables with the constraint that these components maximize the 151 covariance with the response variables (Martens, 2001). In its symmetric form, CoCA fits 152 in the framework of co-inertia analysis (COIA, Dolédec & Chessel, 1994) that is not 153 affected by the problem of collinearity (Dray, Chessel, & Thioulouse, 2003). Co-inertia 154 analysis relates two data tables in a symmetric way, by providing a decomposition of the 155 co-inertia criterion on a set of orthogonal axes on which sample scores are projected 156 (Dray, Chessel, & Thioulouse, 2003).

157 In practice, CoCA identifies associations (or common ecological gradients) 158 between two types of biological assemblages from the same samples, by seeking the 159 factorial axes that maximize the covariance between the weighted average sample scores (projection of rows of one table onto the factorial axes) of one community with 160 161 those of the other community (Supporting Information, Eq. 12). This requires to determine species scores (projection of columns of one table onto the factorial axes) of 162 one table as weighted averages of sample scores of the other table and sample scores as 163 164 weighted averages of the species scores of their own table (Supporting Information, Eq. 165 <u>13–16</u>). So, pCoCA and sCoCA are a weighted version of PLS and COIA, respectively. 166 Given that <u>CoCA is related to the correspondence analysis</u>, it is necessary to circumvent the fact that the sample weights of Y and X (R_1 and R_2 , which are the row sums of Y and 167 **X**, respectively) are imposed and are not similar (Supporting Information, Eq. 6 and Eq. 168 169 7). This does not meet a crucial constraint in estimating the co-structure between two tables, namely that the samples must be weighted in the same way for the two tables 170

171 (Dray, Chessel & Thioulouse, 2003). Hence, an additional common sample weights

- 172 matrix (R_0) is defined to replace R_1 and R_2 so that the weighted averaging properties of
- 173 CA can be retained in CoCA (ter Braak & Schaffers, 2004). In pCoCA, R_0 is equal to the
- 174 sample weights of the community response table (i.e., $R_0 = R_1$), whereas in sCoCA,
- 175 $R_0 = (R_1 + R_2)/2$ (Supporting Information, Eq. 10 and Eq. 11). Note that in the specific case

176 of compositional data, the row sums are equal to 1 and $R_0 = R_1 = R_2$.

177

178 Ordination of the structure and assemblage of interacting communities

179 From sCoCA, ordination diagrams can be made in the usual way by jointly plotting the species scores and sample scores (of each community) for the first axes of 180 181 the analysis (ter Braak & Schaffers, 2004). For an optimal representation of this 182 association in a biplot, the species scores of each axis must be multiplied by the quarter 183 root of the eigenvalue of the axis (ter Braak, 1990). From pCoCA, the fit of the response 184 community (viruses) to the predictive community (microalgal hosts) as well as the 185 variations in the composition of communities can be displayed in ordination diagrams 186 (biplots; ter Braak & Schaffers, 2004). For instance, the joint plot of sample scores of 187 hosts (table X) with species scores of viruses (table Y) displays the fit of the virus OTUs 188 from the host communities. The simultaneous plot of sample scores of microalgae with 189 the loadings (i.e., the coefficient or importance of variables on the first components of 190 the PLS) of predictor species (i.e., hosts) allows representing microalgae communities and their OTUs. Both types of OTUs can also be jointly displayed using scores of the 191 192 response species (i.e., viruses) with the loadings of predictor species (i.e., microalgae). 193 For all these diagrams, axes are optimized to maximize covariance between assemblages 194 rather than to depict associations within individual species matrices, but interpretation

can also be carried out as in correspondence analysis according to the barycentre

196 principle, where OTUs are placed at the barycentre (weighted average) of the sample

197 points and, by symmetry, samples at the barycentre (weighted average) of the OTUs

198 (Supporting Information, Eq. 13–16).

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200 **Real data application**

We illustrate the use of CoCA to study the congruence between microbial
communities from NGS data, employing data of Simon et al. (2015) on the study of
autecology of microbial eukaryotes in shallow freshwater systems, and a data set
acquired during our research program (ANR program DECOVIR-12-BSV7-0009) on the
monitoring of microalgae and viruses in marine systems.

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207 Case study 1: Microbial eukaryotes

208 Surface water was sampled monthly from April 2011 to April 2013 in five small 209 shallow freshwater systems [four ponds: Etang des Vallées (EV), La Claye (LC), Mare 210 Gabard (GB), Saint Robert (SR), and one brook: Ru Sainte Anne (RSA)], located at the Natural Regional Park of the Chevreuse Valley (South of Paris, France). These systems 211 212 were characterized by different local environmental conditions. Briefly, raw genomic 213 sequences were obtained from 18S rDNA fragments, encompassing the V4 214 hypervariable region, applying 454 pyrosequencing and filtered to remove potential 215 spurious sequences using a local pipeline (Simon et al., 2014). Sequences from all 216 samples were then processed together and clustered into OTUs at 0.98 similarity cutoff 217 using CD hit (Fu, Niu, Zhu, Wu, & Li, 2012), and singletons were eliminated, before to 218 assign OTUs to taxonomic groups based on sequence similarity to the PR2 database 219 (Guillou et al., 2013). From this overall OTU table, we used the method applied by Simon et al. (2015) to select the most abundant OTUs. Finally, we subdivided the data set by grouping OTUs in two functional groups (autotrophic and heterotrophic) trained on literature information (Simon et al., 2015; Genitsaris, Monchy, Breton, Lecuyer, & Christaki, 2016 and references within), and we obtained one table for autotrophic microbial eukaryotes (n=108, p=122) and another table for microbial heterotrophic eukaryotes (n=108, q=104).

226

227 Case study 2: Microalgae-virus system

228 The data set coming from our research program contains a NGS-based eukaryotic 229 microalgae community table (photosynthetic picoeukaryotes in the class 230 Mamiellophyceae) and a NGS-based virus community table (viruses infecting this class 231 of eukaryotic phytoplankton and belonging to the genera *Prasinovirus* of the family 232 Phycodnaviridae). The data have been acquired across four sites located in northwest Mediterranean Sea (Gulf of Lion) and sampled monthly from March 2013 to April 2014. 233 The Gulf of Lion is characterized by contrasted environments, including eutrophic 234 lagoons connected to the sea, nutrient-rich coastal sites, and oligotrophic open-sea 235 236 locations. The sample locations included two sites in Leucate lagoon (one coastal site 237 (LB) and another site (LA) at the level of the Grau, i.e. the connection with the sea), a 238 coastal site (SA, marine station included in the French marine monitoring network 239 SOMLIT) and an open-sea site (MA, marine station included in the monitoring network 240 MOOSE). The characterization of *Prasinovirus* was based on analyzing the partial 241 sequence of the DNA polymerase gene (PolB) amplified using two primer sets (Chen & 242 Suttle, 1995; Clerissi et al., 2014). For Mamiellophyceae, the sequence of the V9 region of 243 the 18S rDNA was amplified using primers defined by Amaral-Zettler, McCliment, 244 Ducklow, & Huse (2009). The genomic sequences of PolB and V9 region were amplified

245 and sequenced using an Illumina MiSeq platform (GeT-PlaGe, INRA, Castanet-Tolosan, 246 France). Sequences were processed and clustered into OTUs at 0.99 similarity cutoff and 247 singletons were removed using MOTHUR v 1.35.1 (Schloss et al., 2009) for 248 Mamiellophyceae and USEARCH v7 (Edgar, 2010) combined with MUSCLE software 249 (Edgar, 2004) for Prasinovirus. Sequences were then compared against the PR2 database 250 (Guillou et al., 2013) and the NCBI database for Mamiellophyceae and *Prasinovirus*, 251 respectively, in order to assign OTUs to taxonomic groups based on similarity. Finally, 252 we focused specifically on OTUs assigned to the family Mamiellales of the class 253 Mamiellophyceae, and notably the genus Bathycoccus, Micromonas and Ostreococcus 254 which usually dominate this class in the Gulf of Lion and more generally the 255 picoeukaryotic fraction in other ecosystems (Wu, Huang, & Zhong, 2013; Zhu, Massana, 256 Not, Marie, & Vaulot, 2005). Subsequently, the Prasinovirus data set was limited to OTUs 257 assigned as Bathycoccus viruses (BpVs), Micromonas viruses (MpVs) and Ostreococcus viruses (OtVs). The dominant microalgae and virus OTUs (i.e., $\geq 0.1\%$ of mean relative 258 259 abundance for at least two samples) were selected to obtained one microalgae table (n=31, p=67) and one virus table (n=31, q=98). 260

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262 Statistical analysis

Tables X and Y of autotrophic and heterotrophic microbial eukaryotes,
respectively, were subjected to sCoCA. <u>To test the significance of the global co-variation</u>
between the two tables, <u>a</u> Monte-Carlo permutation procedure with 9999 permutations.
<u>was used</u>. In each permutation, sCoCA (<u>by considering all axes</u>) is reapplied to obtain a
value of the co-variance between table Y and row-permuted X (so that samples are
randomized while preserving the relative abundance of individuals). Note that the
choice of the table to be reordered is not important here since we used the symmetric

270 form of the CoCA. A null distribution was estimated from co-variance calculated for the 271 permuted data. The observed co-variance is then compared to the distribution obtained 272 under the null hypothesis. The positions of the samples on ordination axis of each table 273 are then correlated to show the overall level of co-variation between them. For the 274 microalgae-virus data set, both tables were subjected to pCoCA with the SIMPLS 275 algorithm. ter Braak & Schaffers (2004) suggest that the number of axes used to 276 summarize the data can be selected by "leave-one-out" cross-validation procedure to 277 maximize the cross-validatory fit (%) that measures how well the table X (microalgae in 278 our case) predicts the response table Y (viruses in our case). Working just with these 279 significant axes provides a measure of association between the tables by removing random noise and keeping only the major dimensions of ecological variability. Finally, 280 281 we combine CoCA and network analysis so that nodes in <u>co-occurrence</u> networks are 282 reordered according to the species scores from the CoCA, and thus from the co-structure between communities. A novel extension of SPIEC-EASI (Tipton et al. 2018) was used to 283 infer the cross-group <u>co-occurrence</u> networks between two data sets. We used the 284 285 neighborhood (MB) setting and selected the optimal sparsity parameter based on the 286 Stability Approach to Regularization Selection (StARS) (Liu, Roeder, & Wasserman, 287 2010). The StARS variability threshold was set to 0.05 for networks built from the two 288 data sets. All statistical analyses were performed with the R software (R Core Team, 289 2019) and using the cocorresp package (Simpson, 2009) for CoCAs and the SpiecEasi 290 package (Kurtz et al., 2015) for <u>co-occurrence</u> networks. <u>Appendix 1 contains a</u> R script_ 291 and example data (from the case study 1 and the case study 2) allowing users to 292 reproduce the analysis and apply them on their own data sets.

- 293
- 294 Results

295 Case study 1: Microbial eukaryotes

The common variance between the two microbial groups computed from the 296 297 sCoCA explained, significantly (), 13.21% of the total variation of autotrophic microbial 298 eukaryotes and 15.22% in heterotrophic microbial eukaryotes. Of the common variance, 299 39.51% was accounted by the first three axes of the sCoCA (sCoCA axis 1: 18.95%, 300 sCoCA axis 2: 10.55%, sCoCA axis 3: 10.01%). The first three ordination axes of the 301 autotrophic eukaryotes were highly correlated with the first three ordination axes of the 302 heterotrophic eukaryotes (correlations being 0.95, 0.92, and 0.92), demonstrating a high 303 degree of similarity in change between autotrophic and heterotrophic microbial 304 eukaryote assemblages. Communities of autotrophic and heterotrophic microbial 305 eukaryotes covaried along a brook/pond gradient on the first axis (from left to right), 306 and an inter-pond variability on the second axis (Figure 1). Marked differences in the 307 composition of the two communities are visible in joint plots. In the brook system (i.e., RSA), the heterotrophic microbial eukaryote community is mainly composed of fungi, 308 309 MAST, Labyrinthulida, and Telonema, whereas in pond systems (i.e., EV, LC, MG, SR) Ciliophora, Biocosoecida, Katablepharida, and Choanoflagellida dominated the 310 communities (Figure 1a). Differences between pond systems are explained with higher 311 312 relative abundance of Biocosoecida, and Katablepharida in EVs and LCs and higher 313 relative abundance of Ciliophera in SRs. Species scores of autotrophic microbial 314 eukaryotes indicated that the patterns in heterotrophic communities are associated to a 315 structure of the autotrophic community (Figure 1b). In the brook system, the 316 autotrophic microbial eukaryote community exhibits high relative abundances of 317 specific OTUs of Bacillariophyceae, Chrysophyte and Cryptophyta. In pond systems, 318 autotrophic communities made up mainly of other specific OTUs of Chlorophyta,

Chrysophyte, and Cryptophyta. It is also worth noting that Dinophyta and Streptophyta
were found exclusively in pond systems (in particular in MG and SR, respectively).

321 The community organization of microbial eukaryotes was highlighted from a 322 cross-group co-occurrence network between autotrophic and heterotrophic individuals. 323 Among 226 dominant autotrophic and heterotrophic OTUs, 204 displayed 274 324 associations (Figure 2). From these associations between OTUs in the network, more 325 positive (98.5%) than negative associations were inferred. All negative associations 326 occurred between OTUs assigned to Ciliphora for heterotrophic microorganisms and 327 Cryptophyta for autotrophic microorganisms. No clear association patterns can be 328 identified in network from raw tables of autotrophic and heterotrophic microbial 329 eukaryotes (Figure 2a). When the co-occurrence network is reordered according to 330 species scores on the first axis of sCoCA, two modules can be distinguished (Figure $\underline{2}b$). 331 The first module (i.e., top right corner) is constituted by OTUs exhibiting the higher 332 relative abundances in brook system (i.e., RSA), while OTUs that compose the second module (i.e., bottom left corner) dominate pond systems (i.e., EV, LC, MG, and SR). 333 334 Heterotrophic OTUs exhibited major associations with autotrophic OTUs belonging to the same module, with only 1.5% of associations between OTUs from distinct module. A 335 336 striking pattern is that Chrysophyte is the autotrophic group that contributes most to 337 associations in the two modules (module 1: 71%, module 2: 41%), whereas for the 338 heterotrophic group it is fungi in the module 1 (47%) and Ciliophora in the module 2 (59%). In pond systems (i.e., module 2), surprisingly, fungi are involved in very few 339 340 associations (7%).

341

342 Case study 2: Microalgae-virus system

343 The cross-validation procedure identified the best pCoCA model based on the 344 first two significant axes (pCoCA axis 1: p=0.001, pCoCA axis 2: p=0.001), in which 345 Mamiellophyceae community predicted 32.02% of the variation in *Prasinovirus* 346 community. The first two axes accounted for 37.47% (24.26% and 13.21% for axis 1 and 347 2 respectively) and 44.82% (28.01% and 16.81%) of the variation in the structure of 348 Mamiellophyceae and *Prasinovirus* community respectively. The biplots indicated that 349 the two communities covaried along a lagoon (samples LAs)/open-sea gradient 350 (samples MAs) on the first axis (from left to right), while a temporal gradient for site LA 351 (intra-site variability) could be identified along the second axis (Figure $\underline{3}$). OtVs have a 352 higher prevalence in the lagoon samples (especially LAs) and coastal samples (SAs) in 353 which *Ostreococcus* exhibited a high density (Figure <u>3</u>a, b). Conversely, open-sea 354 samples (MAs) were dominated by Bathycoccus which supported virus assemblages 355 dominated by BpVs. Micromonas showed a wider distribution, with a relative contribution of its OTUs both in lagoon samples, coastal samples and open-sea samples, 356 357 associated with a similar repartition of MpVs (Figure $\underline{3}a, b$). 358 Based on the cross-group co-occurrence networks analysis, 67 associations were 359 identified between the major 67 OTUs assigned to one of the three groups of 360 Mamiellophyceae (i.e. Bathycoccus, Micromonas, and Ostreococcus) and the major 98 361 OTUs assigned to Prasinovirus (i.e. BpVs, MpVs, and OtVs) (Figure 4a). Reordering the 362 <u>co-occurrence</u> network according to the species scores on the first axis of pCoCA 363 highlighted a structure in the network (Figure 4b). The network topology suggests that 364 the identity of OTUs contained in co-occurring groups of viruses and microalgae are 365 related to their respective prevalence along the lagoon/open-sea gradient. Virus OTUs 366 mostly present in lagoon samples have significant associations primarily with 367 microalgae OTUs displaying the higher prevalence in lagoon samples (top right corner,

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368 Figure 4b). Similarly, virus OTUs dominating the open-sea samples were mainly 369 associated with microalgae OTUs from open-sea samples (bottom left corner, Figure 4b). 370 Among the associations contained in the network, 50 were identified between OTUs 371 belonging to an expected host-virus system (i.e. associations *Bathycoccus*/BpV, 372 Micromonas/MpV, and Ostreococcus/OtV) while the other 17 significant associations 373 were found between OTUs belonging to different host-virus systems. In average 50%, 374 70.6% and 91.3% of associations found for OTUs of BpVs, MpVs, and OtVs respectively 375 were with OTUs assigned to their respective host group. Within the associations, some 376 single Mamiellophyceae OTUs were associated with many Prasinovirus OTUs and 377 reciprocally. On the other hand, dyads were identified, that is specific associations, in 378 Bathycoccus/BpV, Micromonas/MpV and Ostreococcus/OtV systems. Few negative 379 associations inferred from the observation that those OTUs do not co-occur were found in network (Figure 4b). Interestingly, eight negative associations from a total of nine 380 involved virus OTUs and microalgae OTUs belonging to different host-virus systems. 381

382

383 Discussion

384 Critical review and guidance papers on the analysis of NGS-based community 385 data (Ramette, 2007; Buttigieg & Ramette, 2014; Paliy & Shankar, 2016) do not mention 386 any direct quantitative method for predicting the composition of one community from 387 another. Co-correspondence analysis (ter Braak & Schaffers, 2004) fills this gap. At the 388 level of the case study 1, symmetric form of CoCA indicated that heterotrophic microbial 389 eukaryote assemblages in shallow freshwater ecosystems were strongly associated with 390 patterns of autotrophic microbial eukaryotes presence and abundance with links that 391 can be taxon-specific (Figure 1). Our results shown also that the composition of 392 heterotrophic microbial eukaryote community was dominated by fungi in brook system

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393 compared to ponds. This is in line with recent observations, based on the estimation of 394 ergosterol level, of a generally higher fungal biomass in river than in ponds (Baldy et al., 395 2002). Higher fungal abundance might potentially be linked to incoming resources from 396 runoff, since in brooks the most important source of imported material is usually 397 deciduous leaves, whose the decomposition processing involved to a large extent fungi 398 (Bärlocher, 1985; Webster & Benfield, 1986). To this, the composition of heterotrophic 399 microbial eukaryote community characterizing brook system is associated a specific 400 composition of microbial autotrophs. Such a result suggests that heterotroph 401 community composition exert a control on the composition of autotroph community, 402 and that microbial autotrophs can be driver of microbial heterotrophs.

403 <u>Regarding the case study 2, the predictive form of CoCA points out different</u> 404 distribution patterns among the three groups of *Prasinovirus* along the lagoon/open-sea 405 gradient (Figure <u>3</u>). Note the importance of the dimension reduction step in pCoCA that 406 allows focusing on ecological structures depicted on a limited number of axes and removes random variation from the data. The patterns in *Prasinovirus* assemblages, with 407 408 a dominance of OtV in lagoon and coastal samples compared to offshore locations, the 409 inverse distribution for BpV, and MpV exhibiting a wider spatial distribution, are in part 410 a consequence of the presence of their respective hosts in lagoon, coastal and open-sea 411 samples. Indeed, Ostreococcus is known to be abundant in lagoons (Subirana et al., 412 2013), more eutrophic system, compared to *Bathycoccus*, which is found mainly in 413 oligotrophic areas (Vaulot et al., 2012; Wu, Huang, & Zhong, 2013) such as offshore sites 414 (i.e., MA). *Micromonas* is ubiquitous and particularly present in nutrient-rich 415 environments (Not et al., 2004; Viprey, Guillou, Ferréol, & Vaulot, 2008). Our findings 416 confirm also the data of Bellec et al. (2010) showing that OtV are more abundant in 417 lagoon than in the open sea.

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418 Cross-taxon congruence description and evaluation (Virtanen, Ilmonen, Paasivirta, & Muotka, 2009; Gioria, Bacaro, & Feehan, 2011) provides a more 419 420 comprehensive picture of the community similarity than the richness metrics 421 conventionally used (Wolters, Bengtsson, & Zaitsev, 2006; Westgate, Barton, Lane, & 422 Lindenmayer, 2014). Our results reinforce the need to use CoCA to study the cross-taxon 423 congruence in microbial communities from NGS data. This is all the more important 424 because the high co-correspondence between the two functional groups in microbial 425 eukaryote community may be especially informative given the key ecological role of 426 microbial eukaryotes (Caron et al., 2012). In addition, the study of cross-taxon 427 congruence between Mamiellophyceae and *Prasinovirus* is of definite interest in marine 428 ecosystems, warmed by climate change, where the expected gradual shift towards small 429 primary producers could render the role of small eukaryotes more important than they 430 are today (Morán, López-Urrutia, Calvo-Díaz, & Li, 2010). Microbial eukaryotes are 431 recognized as a significant contributor across various geographical locations of picophytoplankton (Worden, Nolan, & Palenik, 2004; Jardillier, Zubkov, Pearman, & 432 433 Scanlan, 2010), which accounts for > 50% of phytoplankton biomass and productivity in marine ecosystems (Maranon et al., 2001; Teira et al., 2005). 434

435 Previous studies have suggested that biotic interactions are the most likely 436 mechanisms underlying cross-taxon congruence at local scales (Jackson & Harvey, 1993; 437 Johnson & Hering, 2010), although concordance is also expected from similar responses 438 to environmental gradient (Bini, Vieira, Machado, & Machado Velho et al., 2007; Rooney 439 & Bayley, 2012). As an important implication, the level of congruence can inform about the structural pattern among interacting groups (Özkan et al., 2014). That being said, 440 441 reordering <u>co-occurrence</u> networks, from the species scores on the first axis of CoCA, 442 allows to identify structural patterns in <u>co-occurrence</u> networks of microbial eukaryote

443 community (Figure $\underline{2}$ b) and in the Mamiellophyceae/*Prasinovirus* system (Figure 4b). 444 For example, the structure of the microbial eukaryote network is characterized by two 445 modules underlying a brook/pond gradient in the composition of heterotrophic and 446 autotrophic microbial eukaryote assemblages. These differences in OTU composition 447 within the two modules suggest that the food-web structure is different between lotic 448 and lentic ecosystems, and reinforce the differences previously observed at the 449 bacterioplankton level (Portillo, Anderson, & Noah, 2012). A substantial effort has been 450 made in the development of metrics to estimate and test the level of nestedness (e.g., 451 Rodriguez-Guirones & Santamaria, 2006; Ulrich & Gotelli, 2007) and modularity (e.g., 452 Barber, 2007; Dorman & Strauss, 2014) within interaction networks. The order of 453 individuals of two groups in bipartite matrices affects the magnitude of metrics that 454 represent deviations from an idealized state (e.g., perfect nestedness or modularity) 455 (Almeida et al., 2008). It has then been advocated that prior to analysis of structure of 456 networks, original bipartite matrices should be reordered to maximize the coherence of 457 individual distributions in rows and columns, so that individuals with most similar links are close together (Borgatti & Everett, 1997; Leibold & Mikkelson, 2002). These findings 458 taken together with our results validate our proposition to combine CoCA with network 459 460 analysis to study structural patterns of microbial networks. In addition, in reverse to the 461 expected view of nestedness structure of phage-bacteria network (Flores, Meyer, 462 Valverde, Farr, & Weitz, 2011), the modular structure of Mamiellophyceae/Prasinovirus 463 network observed in the field underpinned the modularity patterns previously observed 464 in phage-bacteria network from cross-infection experiments (Flores, Valverde, & Weitz, 465 2013). Given that the structure of interaction networks is constraint by the 466 coevolutionary processes between species (Peralta, 2016), this would lead to account

467 for phylogenetic signals within co-occurrence network (Derocles et al., 2018). In this

468 context, it would be possible to disentangle the confounding effect of phylogeny from.
469 true biotic interactions by developing a partial analysis (ter Braak, Šmillauer & Dray,
470 2018) in the context of CoCA to partial out the phylogenetic effect and focus on patterns.
471 of co-occurrence that are not related to phylogenetic signal.

472 All computational methods used to infer networks from NGS data sets produce 473 species <u>co-occurrence</u> networks, where a link between two species represents a 474 significant statistical association (positive or negative) between their abundance. This 475 raises a critical issue about the interpretation of inferred associations (Derocles et al., 476 2018), because <u>co-occurrence</u> networks differ from interaction networks constructed on 477 observations of both the species and their interactions (Ings et al., 2009). For instance, 478 all inferred associations between Mamiellophyceae and Prasinovirus belonging to 479 expected host-virus system were positive. These results are consistent with a previous 480 work showing that when parasitism is captured as a significant link in <u>co-occurrence</u> 481 network, it is retrieved as a positive link despite the detrimental effect of parasite on its host (Weiss et al., 2016). This might be explained because the copresence of the host 482 483 species and the parasite species is necessary for the interaction to occur. Another 484 surprising result is that all negative associations (expected one) were between 485 Mamiellophyceae and *Prasinovirus* belonging to different host-virus systems. Such 486 negative associations may account for opposite abiotic requirements, since in our case, 487 OTUs of concerned viruses and microalgae had inverse spatio-temporal dynamics. 488 Positive associations were also found between individuals of different host-virus 489 systems, which could be explained by the increase of a Prasinovirus population that 490 removes, by infection, a major competitor of a co-occurring Mamiellophyceae host of 491 another group. It is important to keep in mind that, although these associations between 492 microalgae and *Prasinovirus* suggest that they interact, they do not necessarily mean

493 that the co-occurring Mamiellophyceae are the virus hosts, even if they belong to the 494 expected group of hosts (e.g., Bathycoccus, Micromonas or Ostreococcus). Our approach 495 (CoCA combined with network analysis, or other method to infer associations) could be 496 combined with other approaches, e.g. single cell genomics (Kalisky, Baliney, & Quake, 497 2011; Martinez-Garcia et al., 2012), Epic-PCR (Spencer et al., 2016), to validate the 498 predicted associations in interactions. The justification of the association sign between 499 the two functional groups making up the microbial eukaryote community is also not 500 straightforward, although they can be triggered by ecological interactions or by species 501 abiotic requirements (Derocles et al., 2018).

502 In conclusion, the successful application of co-correspondence analysis over two 503 real data sets of microbial communities exhibiting various types of interactions 504 reinforces that resorting to this method for study the distributions, assemblages and 505 interactions between two microbial communities constitutes a highly valuable approach to understand the cross-taxon congruence between microorganisms. A useful 506 consequence of cross-taxon congruence is that the distribution of well-known taxa may 507 provide insight into the processes structuring the distribution of other taxa (e.g., Bilton, 508 509 McAbendorth, Bedford, & Ramsay, 2006; Santi et al., 2010). This approach could be used 510 to enhance our understanding of a major problem, the effect of phytoplankton bloom (in 511 particular toxic groups such as cyanobacteria) on the microbial communities and in turn 512 on the ecosystem functions (e.g., Yang et al., 2016; Xue et al., 2018). Our findings also 513 demonstrate that the reordering of <u>co-occurrence</u> networks, according to the 514 congruence information extracted from CoCA, allows to investigate the presence of 515 ecological signals in networks. The advantage of this approach is that the complexity of 516 the network is considerably reduced by the non-random placement of nodes in the 517 space in such a way as to improve the aesthetic quality of the representation and

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518 consequently its readability, as proposed in good practice of data visualization 519 (Spiegelhalter, Pearson, & Short, 2011; Kjærgaard, 2015; Pocock et al., 2016). Interestingly, the applicability of our approach goes beyond the particular case of data 520 521 sets with row-sum normalisation (i.e., compositional data). Indeed, CoCA was originally designed to analyze abundance data and is thus able to deal with counts without the 522 523 need to rarefy data, in accordance with the recent advice against rarefaction (McMurdie & Holmes, 2014). It paves the way for further studies to examine the cross-taxon 524 525 congruence and structural pattern of <u>co-occurrence</u> networks in microbial communities and in turn their effects on the ecosystem functioning. 526

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821 Data accessibility statement

- 822Data and R scripts to reproduce the different analyses of the case study 1 and the case
- 823 <u>study 2 Mamiellophyceae/Prasinovirus system</u> are available online <u>in the Appendix 1as</u>
 824 <u>additional supporting information</u>.
- 825

826 Author contributions

- 827 Y.D. and S.D. supervised the project; B.A. performed all the statistical analyses; C.J.F.B.
- 828 contributed to mathematical development of the method; H.L. contributed to collecting
- 829 and genotyping the biological materials of microalgae and viruses; B.A. wrote the first
- 830 draft of the manuscript; B.A., C.J.F.B., Y.D., and S.D. commented and approved the final
- 831 version of the manuscript.
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- 833 Supporting information
- 834 Additional supporting information may be found online in the Supporting Information
- 835 section at the end of the article.
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Figures

- Figure 1. Ordination biplots of the case study 1, representing of the communities of the
- positions of sites (open diamond) and species (solid triangle) on the axis $1 \times axis 2$
- factorial plan of the symmetric co-correspondence analysis. (a) Biplot for the
- heterotrophic microbial eukaryotes and (b) biplot for the autotrophic microbial
- eukaryotes OTUs obtained from symmetric co-correspondence analysis. EV: Etang
- des Vallées, LC: La Claye, GB: Mare Gabard, SR: Saint Robert, RSA: Ru Sainte Anne.
- Heterotrophic and autotrophic microbial eukaryotes OTUs are colored according to

the phylogenetic group they belong to. "d" indicates the mesh of the grid.



Figure 2. Heatmap of the case study 1, representing the association network between theof heterotrophic and autotrophic microbial eukaryotes (a) before (a) and (b) after-(b) the reordering the position of each species in the network according to their network from the species scores on the first axis of symmetric co-correspondence analysis. Bar plots of the relative abundance of (c) autotrophic microbial eukaryotes and (d) heteretrophic microbial eukaryotes. Each OTU is represented by a vertical line partitioned into segments corresponding to its relative abundance in one of five sites. EV: Etang des Vallées, LC: La Claye, GB: Mare Gabard, SR: Saint Robert, RSA: Ru Sainte Anne.



863 **Figure 3.** Ordination biplots <u>of, concerning</u> the case study 2, <u>ofrepresenting the</u>

- 864 positions of sites (open diamond) and species (solid triangle) on the axis 1 × axis 2
 865 factorial plan of the predictive co-correspondence analysis. (a) Biplot for the
- 866 Mamiellophyceae and (b) <u>biplot for the</u> *Prasinovirus* OTUs obtained from predictive
- 867 co-correspondence analysis. LB: coastal site in Leucate lagoon, LA: site at the level of
- the Grau in Leucate lagoon, SA: coastal site, MA: open-sea site. Microalgae and virus
- 869 OTUs are colored according to the phylogenetic group they belong to. BpV:
- 870 Bathycoccus viruses, MpV: Micromonas viruses, OtV: Ostreococcus viruses. "d"
- indicates the mesh of the grid.
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and after (b) the reordering the network from the species scores on the first axis of symmetric co-correspondence analysis. Bar plots of the relative abundance of (c)autotrophic microbial eukaryotes and (d) heteretrophic microbial eukaryotes. Each-0TU is represented by a vertical line partitioned into segments corresponding to itsrelative abundance in one of five sites. EV: Etang des Vallées, LC: La Claye, GB: Mare-Gabard, SR: Saint Robert, RSA: Ru Sainte Anne.
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- Figure 4. Heatmap of the case study 2, representing theof microalgae-viruses
 (Mamiellophycaea/Prasinovirus) association network (a) before (a) and (b) after (b)
 the reordering the position of each species in the network according to their network
 from the species scores on the first axis of predictive co-correspondence analysis. Bar
 plots of the relative abundance of (c) microalgae and (d) viruses. Each OTU is
 represented by a vertical line partitioned into segments corresponding to its relative
 abundance in one of four sites. LA: site at the level of the Grau in Leucate lagoon, LB:
- 892 coastal site in Leucate lagoon, SA: coastal site, MA: open-sea site.
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